# **BMJ Open Using primary care data to assess comparative effectiveness and safety of apixaban and rivaroxaban in patients with nonvalvular atrial fibrillation in the UK: an observational cohort study**

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**To cite:** Jaksa A, Gibbs L, Kent S, *et al*. Using primary care data to assess comparative effectiveness and safety of apixaban and rivaroxaban in patients with nonvalvular atrial fibrillation in the UK: an observational cohort study. *BMJ Open* 2022;12:e064662. doi:10.1136/ bmjopen-2022-064662

► Prepublication history and additional supplemental material for this paper are available online. To view these files, please visit the journal online [\(http://dx.doi.org/10.1136/](http://dx.doi.org/10.1136/bmjopen-2022-064662) [bmjopen-2022-064662](http://dx.doi.org/10.1136/bmjopen-2022-064662)).

Received 12 May 2022 Accepted 05 October 2022

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# ABSTRACT

**Objective** To compare real-world effectiveness and safety of direct oral anticoagulants (DOACs) in patients with nonvalvular atrial fibrillation (AFib) for prevention of stroke. Study design and setting A comparative cohort study in UK general practice data from The Health Improvement Network database.

Participants and interventions Before matching, 5655 patients ≥18 years with nonvalvular AFib who initiated at least one DOAC between 1 July 2014 and 31 December 2020 were included. DOACs of interest included apixaban, rivaroxaban, edoxaban and dabigatran, with the primary comparison between apixaban and rivaroxaban. Initiators of DOACs were defined as new users with no record of prescription for any DOAC during 12 months before index date.

Primary and secondary outcome measures The primary outcome was stroke (ischaemic or haemorrhagic). Secondary outcomes included the occurrence of all-cause mortality, myocardial infarction (MI), transient ischaemic attacks (TIA), major bleeding events and a composite angina/MI/stroke (AMS) endpoint.

Results Compared with rivaroxaban, patients initiating apixaban showed similar rates of stroke (HR: 0.93; 95%CI 0.64 to 1.34), all-cause mortality (HR: 1.03; 95%CI 0.87 to 1.22), MI (HR: 0.95; 95%CI 0.54 to 1.68), TIA (HR: 1.03; 95%CI 0.61 to 1.72) and AMS (HR: 0.96; 95%CI 0.72 to 1.27). Apixaban initiators showed lower rates of major bleeding events (HR: 0.60; 95% CI 0.47 to 0.75). Conclusions Among patients with nonvalvular AFib, apixaban was as effective as rivaroxaban in reducing rate of stroke and safer in terms of major bleeding episodes. This head-to-head comparison supports conclusions drawn from indirect comparisons of DOAC trials against warfarin and demonstrates the potential for real-world evidence to fill evidence gaps and reduce uncertainty in both health technology assessment decision-making and clinical guideline development.

# **INTRODUCTION**

In the UK, atrial fibrillation (AFib) affects 1.4 million patients,<sup>1</sup> and between  $0.9\%$  and

# STRENGTHS AND LIMITATIONS OF THIS STUDY

- ⇒ This study performed head-to-head comparisons of direct oral anticoagulants rather than relying on indirect comparisons of trials with different designs.
- ⇒ This study used routinely collected data from electronic health records within a nationally representative UK database with good data recording and follow-up time.
- ⇒ Treatment assignment was not randomised; however, after propensity score matching, treatment groups were similar across 40 measured demographic and clinical characteristics, suggesting comparability between exposure groups.
- ⇒ Robustness of primary findings via propensity score matching was contextualised by use of alternative balancing approaches in sensitivity analyses, including propensity score weighting and high dimensional propensity score matching.
- ⇒ We did not link to secondary data for stroke outcome ascertainment but reporting of stroke and other conditions in the general practice record is incentivised by the Quality and Outcomes Framework.

1.6% of the UK's National Health Service spending is attributable to AFib, predominantly from hospitalisations.[2](#page-10-1) The condition is associated with significant complications, including stroke**—**nonvalvular AFib increases an individual's risk of stroke by five times, $3$ and between 20% and 30% of stroke cases are attributed to AFib.[4](#page-10-3)

Anticoagulants, including vitamin K antagonists (VKAs) and direct oral anticoagulants (DOACs), are highly effective in the prevention and treatment of thromboembolic events associated with AFib. Nevertheless, VKAs like warfarin require frequent coagulation monitoring due to their narrow therapeutic index and have multiple drug and food interactions.

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Alternatively, DOACs inhibit coagulation via direct and specific binding to active sites of thrombin (eg, dabigatran) or factor Xa (eg, apixaban, rivaroxaban and edoxaban) of the coagulation pathway. Compared with VKAs, DOACs have a wider therapeutic index, which permits use in fixed doses without coagulation monitoring, and relatively limited drug and food interactions. DOACs are the preferred anti-coagulants for patients with nonval-vular AFib in the UK.<sup>[5](#page-10-4)</sup> The safety and efficacy of DOACs compared with VKAs for stroke prevention in patients with AFib have been established in randomised clinical trials  $(RCTs)$ .<sup>67</sup> However, head-to-head RCTs of DOACs (eg, apixaban vs rivaroxaban) are not available, and relative safety and efficacy findings are based on indirect comparisons from network meta-analyses  $(NMA)$ .<sup>[7–9](#page-10-6)</sup>

Agencies, like the National Institute for Health and Care Excellence (NICE), evaluate the comparative effectiveness and cost-effectiveness of therapies to inform both reimbursement decisions and clinical guidelines. Between March 2012 and September 2015, NICE separately assessed and recommended four DOACs for stroke prevention in AFib: dabigatran, rivaroxaban, apixaban and edoxaban. Because no direct comparisons of DOACs were available, NICE's decision-making, which impacts patient health and clinical practice was based on RCTs of DOACs compared with warfarin and indirect comparisons to other DOACs. A NMA of indirect comparisons ranked rivaroxaban as the best DOAC for reducing myocardial infarction (MI) and all-cause mortality, while apixaban was ranked best for minimising the risk of bleeding and dabigatran was ranked best for reducing the rate of stroke.<sup>5</sup> However, the NMA rankings had probabilities that varied from 60% to 80% and most of the head-to-head ORs approached the null and/ or had wide confidence intervals leading NICE to interpret the findings with caution. NICE noted heterogeneity among the trials on which the indirect comparisons were based, which limited the ability to differentiate between DOACs' effectiveness. These uncertainties were reflected in NICE's Atrial fibrillation: diagnosis and management guideline<sup>[5](#page-10-4)</sup> where NICE decided to not recommend one DOAC over the others, but instead emphasised treatment should be personalised based on the patients' needs and preference.

Comparative effectiveness analysis in real-world data (RWD) has emerged as a potential strategy for supplementing clinical trials and for generating evidence on the effectiveness of products after launch.<sup>10</sup> There is a growing body of literature that has duplicated AFib RCT results in RWD,<sup>11 12</sup> which increases confidence in RWD studies that directly compare DOACs in RWD. However, there is heterogeneity in the results of RWE studies that directly compare DOACs. For example, a Scottish study in AFib patients found no differences between DOACs for stroke prevention.<sup>13</sup> These findings align with a French and Danish RWE study in nonvalvular AFib patients $^{14}$ <sup>15</sup> but differ from a US-based study of Medicare patients with AFib, which found an increased risk of stroke for

rivaroxaban patients compared with apixaban patients.<sup>[16](#page-10-11)</sup> All-cause mortality findings are also mixed, with one study finding no difference between apixaban compared with rivaroxaban $^{14}$  and others finding increased mortality with rivaroxaban compared with apixaban $^{13}$ <sup>16</sup> and dabigatran.[15](#page-10-12) Only the French and Danish studies restricted to nonvalvular AFib<sup>[14 15](#page-10-10)</sup> and the US-based study<sup>16</sup> was in patients >65 years old. Thus, it is unclear how generalisable these findings are to patients with nonvalvular AFib The objective of this study was to evaluate the comparative effectiveness and safety of the DOACs available in the UK (apixaban, rivaroxaban, edoxaban and dabigatran) through direct comparisons among adults with nonval-We conducted an RWD cohort study to compare the rate of stroke among patients with nonvalvular AFib initiating DOACs, specifically, apixaban, rivaroxaban, edoxaban and dabigatran. Following the steps outlined by Gatto *et*   $a<sup>17</sup>$  $a<sup>17</sup>$  $a<sup>17</sup>$ , we articulated the research question, conceptualised the underlying hypothetical target trial, $^{18}$  $^{18}$  $^{18}$  identified a fitfor-purpose data source, and posted the final protocol publicly on the EU PASS Register of the European Network of Centres for Pharmacoepidemiology and Phar-

# Data source

study.

in the UK.

**METHODS** 

vular AFib at risk for stroke in the UK.

Study design and objectives

This study used anonymised patient data from The Health Improvement Network (THIN) Database (A Cegedim Proprietary Database). The THIN Database is a primary care research database containing anonymised electronic health record data from around 850 UK general practices using the VISION clinical system (since 1994) and contains records for around 20million patients. The THIN database has been well described, and the quality of data collection has been documented in multiple studies.<sup>[19 20](#page-10-15)</sup> THIN has also been shown to be representative of the UK population with respect to demographics, major condition prevalence and mortality rates.<sup>2</sup>

macovigilance ([EUPAS45073\)](https://www.encepp.eu/encepp/viewResource.htm?id=45109) prior to implementing the

### Study population and treatment

We identified adults  $(\geq 18$  years) with nonvalvular AFib newly initiating apixaban, rivaroxaban, edoxaban or dabigatran (with a 365-day washout for any prior DOAC use) between July 2014 and December 2020 ([figure](#page-2-0) 1). Index date was assigned as the date of first qualifying treatment initiation. Patients were required to have at least one medical encounter in the 180 days prior to study index date, be at risk of stroke (general practitioner (GP) assessed CHA2DS2 VASc >1 for men and >2 for women), have no recorded history of study outcomes and have no prior diagnosis of cardiac valve disease, deep vein thrombosis, pulmonary embolism, angina or congenital heart



<span id="page-2-0"></span>Figure 1 Study population for completed comparisons. DOACs, direct oral anticoagulants; GP, general practitioner; MI, myocardial infarction; PS, propensity score; TIA, transient ischaemic attack.

disease. History of congenital heart disease was an exclusion criterion because it is linked to valvular heart disease. Exposures of interest were defined by corresponding Anatomical Therapeutic Chemical (ATC) codes [\(online](https://dx.doi.org/10.1136/bmjopen-2022-064662) [supplemental figure 1;](https://dx.doi.org/10.1136/bmjopen-2022-064662) [online supplemental files 1 and](https://dx.doi.org/10.1136/bmjopen-2022-064662) [2\)](https://dx.doi.org/10.1136/bmjopen-2022-064662).

The primary comparison of interest was apixaban initiators versus rivaroxaban initiators, the two most commonly prescribed DOACs in the  $UK<sup>22</sup>$  Additional comparisons considered and completed are detailed in [table](#page-2-1) 1. Patients with qualifying DOAC prescriptions from

both medications or medication groups being compared (<0.5% of patients with qualifying prescriptions) were indexed according to the chronologically first qualifying prescription.

In the primary comparison, subgroup analyses were conducted by age group  $\langle$  <75 years vs  $\geq$ 75 years), by CHA<sub>2</sub>DS<sub>2</sub> VASc score (0-1, 2-3, 24), by gender, and among patients with vs without each of the following: concomitant aspirin use, prior warfarin use, diabetes mellitus and heart failure.

<span id="page-2-1"></span>

\*Did not meet the sample size criteria of greater than 500 patients per treatment group, and thus inferential analyses were not completed. DOACs, direct oral anticoagulants.

### **Outcomes**

The primary outcome was stroke (ischaemic or haemorrhagic). Secondary outcomes included all-cause mortality, MI, transient ischaemic attacks (TIA), major bleeding events and a composite of angina/MI/stroke (AMS). Major bleeding was defined as a composite of major intracranial (including haemorrhagic stroke), gastrointestinal and urogenital bleeds. Study outcomes were defined by corresponding Read Medical Codes and mapped International Classification of Disease, 10th version (ICD-10) diagnosis codes at the primary care setting [\(online](https://dx.doi.org/10.1136/bmjopen-2022-064662) [supplemental files 1 and 2](https://dx.doi.org/10.1136/bmjopen-2022-064662)).

# **Covariates**

Covariates included in the propensity score (PS) model (described below) included age, gender,  $\text{CHA}_2\text{DS}_2$  VASc score, $^{23}$  year of treatment initiation and history of a number of diagnoses and treatments (see list in [table](#page-4-0) 2). CHA<sup>2DS2</sup> VASc score was estimated using patient history ([online supplemental file 2](https://dx.doi.org/10.1136/bmjopen-2022-064662)) because the GP-assessed CHA<sup>2</sup>DS<sup>2</sup> VAS<sub>c</sub> was not available at time of analysis due to data availability constraints. All covariates were determined based on the literature and clinical knowledge regarding their relationship to the primary and secondary outcomes of interest.

Medication use was identified by ATC codes and assessed in the 12 months prior to and including the index date. Comorbidities were identified by Read Medical Codes and mapped ICD-10 diagnosis codes and were assessed over all available data prior to and including index date. Patient demographics were measured on the index date. High missingness (34%–98% missing; [online supple](https://dx.doi.org/10.1136/bmjopen-2022-064662)[mental table 1](https://dx.doi.org/10.1136/bmjopen-2022-064662)) in diagnostic assessments meant that we dropped some covariates that we planned to include in the PS models. This included marital status, cigarettes per day, alcohol glasses per day, body mass index, hemoglobin A1c (HbA1c), international normalised ratio, glomerular filtration rate and creatinine clearance as measures of renal function. Remaining covariates were assessed dichotomously based on the presence or absence of diagnostic or medication codes in the patient history. Patients were assumed to have had a diagnosis or prescription of the relevant code(s) was found among their records. Otherwise, it is assumed that the patient did not experience the event or was not prescribed the medication, thus resulting in no missing data for these variables. Additional dichotomous characteristics considered for inclusion in the PS model (eg, sepsis) capturing few  $\langle 1\% \rangle$ exposed or unexposed patients were not included in analytic models. See [online supplemental file 2](https://dx.doi.org/10.1136/bmjopen-2022-064662) for definitions of all covariates included in PS models.

The 1-year baseline period specified in the study protocol [\(EUPAS45073](https://www.encepp.eu/encepp/viewResource.htm?id=45109)) for capture of baseline comorbidities was expanded to all prior available data after an observed under-capture of comorbid conditions (eg, hypertension) in baseline using the 1-year baseline period. Results corresponding to the protocol-specified

1-year baseline are reported in [online supplemental](https://dx.doi.org/10.1136/bmjopen-2022-064662)  [tables 2 and 3.](https://dx.doi.org/10.1136/bmjopen-2022-064662)

# PS matching

We used PS matching between exposure groups using 1:1 nearest neighbour matching without replacement (± a calliper of 0.01 of the PS). The PS model included a priori selected covariates assessed prior to treatment index, accounting for over-fitting, positivity violations and covariate instability.[24](#page-10-19)

# Diagnostic phase

In order to progress to the inferential analysis phase, each primary and secondary comparison had to pass a series of diagnostic checks (masked to treatment specific outcomes) including: positivity of variables, baseline confounder balance (an absolute standardised difference (ASD)  $\leq$  0.1),<sup>25</sup> sufficient population-level persistence on treatment (median persistence at least 1year), and confirmation that models were not overfit (≥12 exposed patients per covariate). Adequacy of sample size was also assessed, but insufficient sample size per pre-specified power requirements did not preclude estimation, $26$  as long as each PS matched comparator group had at least 500 patients. Comparisons passing all diagnostic criteria included apixaban versus rivaroxaban, apixaban versus other DOACs (rivaroxaban, edoxaban, dabigatran) and rivaroxaban versus other DOACs (apixaban, edoxaban, dabigatran).

# Pre-specified inferential analysis

In the inferential phase, we executed Cox proportional hazards regression models to estimate HR and 95% CI after PS matching in the overall cohort and within each subgroup of interest. The incidence of stroke and secondary outcomes were assessed for the primary comparison of apixaban vs rivaroxaban. Incidence of stroke was compared for all secondary comparisons. Patients were followed in an 'intention to treat' (ITT) approach starting from the day after their index date. In analyses of stroke, MI, TIA, major bleeding events and AMS, patients were followed until the end of the study period (December 2020) or the first occurrence of the outcome, death or date of last contact with GP. In the analysis of all-cause mortality, follow-up was extended beyond the date of last contact to the date of a death if the death occurred within 90 days of a patient's last contact in order to capture deaths reported after a patient's last contact with their GP. Secondary analysis of the primary outcome in the apixaban versus rivaroxaban comparison included an 'as-treated' approach where in addition to the ITT censoring criteria, patients were censored on termination of exposure, crossover of exposure group, or addition of another DOAC. The exposure termination date was defined as the end of the last continuous prescription (allowing for up to 30-day gaps between end of previous and start of next prescription) plus a 30-day risk window.



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<span id="page-4-0"></span>5







### Pre-specified sensitivity analysis

To evaluate the robustness of our findings, highdimensional PS (HdPS) analysis was used to estimate the association between treatment with DOACs and the primary and secondary outcomes. The HdPS approach is a seven-step algorithm that empirically identifies a pool of covariates from different data dimensions (eg, diagnoses, procedures, medications) based on their prevalence and then selects a subset of the covariates for inclusion in a PS model based on their potential to bias the exposureoutcome association[.27](#page-10-22)

### Post-hoc analysis

Additional post-hoc analyses were conducted to further contextualise the study findings. Glaucoma, a condition not impacted by DOAC use, was assessed as a negative control outcome to assess the possibility of residual confounding after PS matching[.28](#page-11-0) Inverse probability of treatment weighting (IPTW) and standardised mortality ratio weighting (SMR) methods were used to evaluate potential treatment effect heterogeneity in estimates of effect for primary and secondary outcomes and the robustness of results.[29](#page-11-1) Primary analysis findings were additionally assessed by gender.

All data analyses were conducted using Aetion Evidence Platform V.4.45 (2021), software for RWD analysis. Aetion, Inc. [https://www.aetion.com.](https://www.aetion.com)

### Patient and public involvement

There was no patient or public involvement in this study, including development of the research questions, selection of outcome measures, study design, conduct or dissemination of findings.

### **RESULTS**

Out of a total of 5655 patients with new use of at least one DOAC before PS matching ([figure](#page-2-0) 1), 2801 initiated apixaban, 2221 initiated rivaroxaban, 398 initiated edoxaban and 261 initiated dabigatran and met the criteria for inclusion [\(table](#page-4-0) 2).

### Primary comparison: apixaban versus rivaroxaban

A total of 2221 rivaroxaban and 2801 apixaban patients were eligible for inclusion in PS-matched groups. Before PS matching, the apixaban group was more likely to initiate treatment after 2017 and more likely to be woman compared with rivaroxaban patients [\(table](#page-4-0) 2). After 1:1 PS matching, 1839 patients with apixaban and 1839 patients with rivaroxaban were identified ([figure](#page-2-0) 1). Differences in covariate prevalence were minimal, with ASD below 0.1 for all characteristics [\(table](#page-4-0) 2). Due to sample size constraints, stratified analyses by dosage were not completed; dosages of index prescriptions are reported in [online supplemental table 1.](https://dx.doi.org/10.1136/bmjopen-2022-064662) Median follow-up time in the ITT analysis was 845 days [IQR 340, 1368] in the apixaban group and 779 days [322, 1284] in the rivaroxaban group ([online supplemental table 4](https://dx.doi.org/10.1136/bmjopen-2022-064662)). In the 'as-treated'

approach, median follow-up time was shorter compared with ITT and longer in apixaban compared with rivaroxaban (506 days vs 412 days). Rivaroxaban patients were more likely to be censored for switching to another DOAC compared with apixaban patients [\(online supple](https://dx.doi.org/10.1136/bmjopen-2022-064662)[mental table 4](https://dx.doi.org/10.1136/bmjopen-2022-064662)).

In the ITT analysis, the rate of stroke per 1000 personyears was 12.47 in the apixaban group and 13.48 in the rivaroxaban group ([table](#page-7-0) 3). Compared with rivaroxaban, patients initiating apixaban showed similar rates of stroke (HR: 0.93; 95%CI 0.64 to 1.34; [table](#page-7-0) 3, [figure](#page-8-0) 2). Rates of secondary outcomes were also similar between apixaban versus rivaroxaban initiators: all-cause mortality (HR: 1.03; 95%CI 0.87 to 1.22), MI (HR: 0.95; 95%CI 0.54 to 1.68), TIA (HR: 1.03; 95%CI 0.61 to 1.72) and AMS (HR: 0.96;  $95\%$  CI0.72 to 1.27). Apixaban initiators showed lower rates of major bleeding events (HR: 0.60; 95% CI 0.47 to 0.75; [table](#page-7-0) 3, [figure](#page-8-0) 2). In secondary as-treated analysis, apixaban and rivaroxaban were similarly equivalent with respect to rates of stroke (HR: 0.95; 95% CI 0.59, 1.55; [table](#page-9-0) 4).

In subgroup analyses, effects on stroke varied by concomitant aspirin use, history of warfarin use, and  $\mathrm{CHA}_{2}\mathrm{DS}_{2}$ VASc score, though there was limited power to detect statistically significant differences ([table](#page-7-0) 3). In HdPSmatched analyses of stroke and secondary outcomes, results were generally similar or trended toward lower outcome rates among patients with apixaban relative to PS-matched results [\(table](#page-9-0) 4, [figure](#page-8-0) 2). Findings were also similar in post-hoc IPTW and SMR weighted analyses of primary and secondary outcomes [\(table](#page-9-0) 4, [figure](#page-8-0) 2).

In ITT analysis of glaucoma as a negative control outcome, PS-matched and HdPS-matched rates of glaucoma were higher among patients with apixaban, though confidence intervals included the null (PS-matched: HR, 95%CI 1.22, 0.76 to 1.97; HdPS-matched: 1.48, 0.84 to 2.64).

### Secondary comparison: apixaban versus other DOACs

After PS matching, 2276 patients with apixaban and 2276 patients with DOACs other than apixaban (rivaroxaban, edoxaban or dabigatran) were included ([figure](#page-2-0) 1). Treatment groups were balanced after PS-matching, defined as ASD below 0.1 for all characteristics [\(table](#page-4-0) 2). Median follow-up time was longer in the apixaban group (median (IQR) 742 days [298, 1,259]) compared with the other DOACs group (681 days [296, 1170]) [\(online supple](https://dx.doi.org/10.1136/bmjopen-2022-064662)[mental table 4](https://dx.doi.org/10.1136/bmjopen-2022-064662)). Similar rates of stroke were observed between groups (HR: 0.90; 95%CI 0.64 to 1.27; [table](#page-7-0) 3).

### Secondary comparison: rivaroxaban versus other DOACs

After PS matching, 1985 patients with rivaroxaban and 1985 patients with DOACs other than rivaroxaban (apixaban, edoxaban, dabigatran) were included [\(figure](#page-2-0) 1). Treatment groups were balanced after PS-matching, defined as ASD below 0.1 for all characteristics [\(table](#page-4-0) 2). Median follow-up time was shorter in the rivaroxaban group (median (IQR) 784 days [318, 1296]) compared <span id="page-7-0"></span>Table 3 HR of stroke and secondary outcomes among patients with nonvalvular atrial fibrillation newly initiating DOACs after propensity score matching: primary analyses



PS model accounts for age, gender, CHA2DS2 VASc score, year of treatment initiation and the following diagnoses and treatments in baseline: non-major bleeding events, anaemia, diabetes, hypertension, heart failure, osteoporosis/hip fracture, malignant neoplasm, acute kidney injury, chronic kidney disease, asthma/COPD, dementia, aspirin, antiplatelets other than aspirin, warfarin, antimeric preparations, NSAIDs, opioids, SSRIs, antidepressants other than SSRIs, antiepileptics, antipsychotics, benzodiazepines, lipid lowering drugs, insulin, antihyperglycemics other than insulins, antihypertensives, antiarrhythmics, nitrates cardiac vasodilators, cardiac stimulants, gastrointestinal protective agents, bisphosphonates and other agents affecting bone structure, systemic corticosteroids, antineoplastics, systemic antibiotics, systemic antivirals, vaccines/immunoglobulins.

DOACs comprised apixaban, rivaroxaban, edoxaban, dabigatran.

In analysis of stroke, MI, TIA, major bleeding events, and AMS, patients were followed until occurrence of outcome, death, end of patient registration or end of study period (December 2020).

\*In analysis of all-cause mortality, patients were followed until occurrence of outcome (death), end of study period (December/2020) or the later date of end of patient registration and any recorded death within 90 days of end of patient registration. Propensity score matched sample size for analysis of all-cause mortality differs from sample size in analysis of other outcomes because of differences in censoring criteria, which impact a small number of patients' eligibility for inclusion in analysis at the start of follow-up. aHR, adjusted HR; PS, propensity score; PY, personyears.



<span id="page-8-0"></span>Figure 2 Hazard ratios of stroke and secondary outcomes, apixaban versus rivaroxaban, intent to treat analyses. PS model accounts for age, gender, CHA2DS2 VASc score, year of treatment initiation and the following diagnoses and treatments in baseline: non-major bleeding events, anaemia, diabetes, hypertension, heart failure, osteoporosis/hip fracture, malignant neoplasm, acute kidney injury, chronic kidney disease, asthma/chronic obstructive pulmonary disease, dementia, aspirin, antiplatelets other than aspirin, warfarin, antianemic preparations, non-steroidal anti-inflammatory drugs (NSAIDs), opioids, selective serotonic reuptake inhibitors (SSRIs), antidepressants other than SSRIs, antiepileptics, antipsychotics, benzodiazepines, lipid lowering drugs, insulin, antihyperglycemics other than insulins, antihypertensives, antiarrhythmics, nitrates cardiac vasodilators, cardiac stimulants, gastrointestinal protective agents, bisphosphonates and other agents affecting bone structure, systemic corticosteroids, antineoplastics, systemic antibiotics, systemic antivirals, vaccines/immunoglobulins Patients were followed until occurrence of outcome, death, end of patient registration, or end of study period (December 2020). HdPS, high dimensional propensity score; IPTW, inverse probability of treatment weighting; PS, propensity score; SMR, standardised mortality ratio weighting.

with the DOACs other than rivaroxaban (828 days [343, 1399])[\(online supplemental table 4](https://dx.doi.org/10.1136/bmjopen-2022-064662)). There was no difference in rate of stroke between groups (HR: 0.96; 95%CI 0.67 to 1.36; [table](#page-7-0) 3).

### **DISCUSSION**

Among patients with nonvalvular AFib, apixaban was as effective as rivaroxaban in reducing the rate of stroke, allcause mortality, MI, TIA and AMS, and was safer in terms of major bleeding episodes. Because there was heterogeneity across prior study results, our results align with some but not all prior research. The Scottish, Danish, and French studies also found similar effectiveness between apixaban and rivaroxaban, $13-15$  while the US study found apixaban to be more effective.[16](#page-10-11) Our effectiveness findings are also consistent with the indirect comparison of apixaban and rivaroxaban in the NMA used in NICE's 2021 clinical guidelines.<sup>[5](#page-10-4)</sup> Prior research has consistently found apix-aban to be safer in terms of major bleeding.<sup>[5 13 14 16](#page-10-4)</sup> Some prior research found similar equivalencies with respect to all-cause mortality,  $5^{14}$  while others have found increased risk of all-cause mortality for rivaroxaban compared with apixaban.<sup>13 16</sup> In the secondary comparisons, the rate of stroke was similar between apixaban and other DOACs and

rivaroxaban and other DOACs. These secondary comparisons offer an increase in referent group sample size with the inclusion of patients with edoxaban and dabigatran, noting that referent groups primarily comprised patients with rivaroxaban or apixaban. Apixaban and rivaroxaban were initiated more often than dabigatran and edoxaban, which is consistent with prescribing patterns in the  $UK<sup>30</sup>$  and aligns with the selection of the primary and secondary comparisons in the study. Because dabigatran and edoxaban were seldom used and sample sizes were small, head-to-head comparisons including these DOACs were not performed.

While RCTs demonstrate that DOACs are non-inferior to warfarin for stroke prevention in patients with nonvalvular AFib, $31-34$  there is a lack of RCT head-to-head comparisons of individual DOACs. This uncertainty can make reimbursement, clinical decision-making and clinical guideline development challenging. NICE recently updated its AFib clinical guidelines, $5$  relying on a NMA and indirect comparisons of individual DOACs for evidence on their comparative effectiveness and safety. This study addresses a known evidence gap and adds to the clinical evidence base from RCTs and NMAs to directly compare DOACs for stroke prevention in patients with

<span id="page-9-0"></span>Table 4 HR of stroke and secondary outcomes among patients with nonvalvular atrial fibrillation newly initiating apixaban versus rivaroxaban after propensity score matching: sensitivity analyses



PS model accounts for age, gender, CHA2DS2 VASc score, year of treatment initiation and the following diagnoses and treatments in baseline: non-major bleeding events, anaemia, diabetes, hypertension, heart failure, osteoporosis/hip fracture, malignant neoplasm, acute kidney injury, chronic kidney disease, asthma/COPD, dementia, aspirin, antiplatelets other than aspirin, warfarin, antiameric preparations, NSAIDs, opioids, SSRIs, antidepressants other than SSRIs, antiepileptics, antipsychotics, benzodiazepines, lipid lowering drugs, insulin, antihyperglycemics other than insulins, antihypertensives, antiarrhythmics, nitrates cardiac vasodilators, cardiac stimulants, gastrointestina protective agents, bisphosphonates and other agents affecting bone structure, systemic corticosteroids, antieoplastics, systemic antibiotics, systemic antivirals, vaccines/immunoglobulins. In assessment of stroke, MI, TIA, major bleeding events and AMS, patients were followed until occurrence of outcome, death, end of patient registration or end of study period (December 2020).

In assessment of all-cause mortality, patients were followed until occurrence of outcome (death), end of study period (December 2020) or the later date of end of patient registration and any recorded death within 90 days of end of patient registration.

HdPS, high dimensional propensity score; IPTW, inverse probability of treatment weighting; PS, propensity score; SMR, standardised mortality ratio weighting.

nonvalvular AFib in the UK. This additional evidence can help agencies like NICE contextualise the comparative effects of rivaroxaban and apixaban and inform its clinical decision-making.

This study has several strengths. The THIN data contain general practice records for over 20million patient records in the UK and are likely reflective of typical patterns of treatment and care in the UK. Prescribing data are comprehensive and complete and are captured prospectively before outcome events. Nevertheless, this study is subject to several limitations. We did not link to secondary care data for stroke outcome ascertainment. However, stroke is included as part of the Quality and Outcomes Framework, $35$  where general practices are incentivised to comprehensively document instances of stroke occurring in other care settings within patients' medical record. THIN data have been used and validated in published literature to evaluate AFib and stroke.<sup>19 36-39</sup> We were not sufficiently powered for any of the comparisons according to the protocol power guidelines. However, an underpowered study still provides valuable information on clinical outcomes. Indeed, Hernan argues that observational causal inference studies, which place little burden on patients, should proceed even if underpowered so that the evidence can be combined with that from other studies through meta-analysis. $26$  For the evaluation of economic outcomes, that is, through a cost-effectiveness analysis, uncertainty in input parameter

values is propagated through the use of probabilistic sensitivity analysis and thus properly reflected in the outputs. Our analysis of glaucoma as a negative control outcome indicates the possibility of residual confounding, with ITT estimates showing a higher rate of glaucoma for apixaban relative to rivaroxaban. Despite possible residual confounding resulting in bias away from the null, we still observed null findings across primary and sensitivity analyses in our primary outcome, stroke. As with the primary analysis, sample size was also a limitation of these sensitivity analyses.

THIN data contain records of prescriptions written, but it is not known whether medications were dispensed or taken. However, most patients initiating apixaban (80.9%) and rivaroxaban (76.8%) had a second prescription within 30 days of the end of their index prescription and the proportion of days covered in the follow-up period was high in both groups (mean 88.1% apixaban and mean 83.4% rivaroxaban), providing evidence that medications were being taken and refilled. Due to data availability constraints, in our analyses, we estimated patient CHA2DS2 VASc scores using patient history, which may deviate from a GP assessment-based score. However, 99.7% of female patients and 99.8% of male patients had patient history-estimated CHA₂DS₂ VASc scores, which aligned with the GP assessment-based score cutoffs imposed by the data vendor ( $\geq 2$  and  $\geq 1$ , respectively).

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This study showed that among patients with nonvalvular AFib, apixaban was as effective as rivaroxaban in reducing rate of stroke and safer in terms of bleeding adverse events. This study demonstrates that comparative effectiveness RWE studies have the potential to fill evidence gaps and reduce uncertainty in HTA decisionmaking and clinical guideline development.

Correction notice This article has been corrected since it was published online. The year has been changed from "2022" to "2021" in the Ethics statement.

Contributors AJ, SK, SR, SD, PJ and NG contributed to the conception and design of the study. AJ was responsible for the acquisition of the data. LG, AA, AP and PG carried out data analysis. AJ, LG, AA, AP and NG interpreted the results. AJ and LG drafted the manuscript. SK, SR, SD, MS, LK, AA, AP, PG, PJ and NG offered critical revisions and gave final sign off on the manuscript. AJ acted as the guarantor.

Funding The authors have not declared a specific grant for this research from any funding agency in the public, commercial or not-for-profit sectors.

Competing interests AJ, LG, AA, AP and NG are employed by and have ownership stake and/or hold stock options in Aetion. Inc. NG holds stock in Pfizer Inc. SK, SR, SD, MS, LK and PJ are employees of the National Institute for Health and Care Excellence. PJ is on the board of the GetReal Institute, which receives grants from Innovative Medicines and Horizon Europe.

Patient and public involvement Patients and/or the public were not involved in the design, or conduct, or reporting, or dissemination plans of this research.

Patient consent for publication Not applicable.

Ethics approval The study concept was approved by the THIN Scientific Review Committee on 6 July 2021. Protocol number 21-013.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement No data are available.

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### Non-Interventional Study Protocol





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# <span id="page-15-0"></span>**2. List of Abbreviations**



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# <span id="page-17-0"></span>**3. Amendments and Updates**



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# <span id="page-18-0"></span>**4. Milestones**



\* The interim report will comprise descriptive and primary objective results for primary comparison.

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# <span id="page-19-0"></span>**5. Rationale and Background**

Health Technology Assessment (HTA) agencies, like the National Institute for Health and Care Excellence (NICE), are turning toward coverage with evidence development and/or managed access programs to allow additional time for data on the effectiveness of products to mature or be collected. Real-world data (RWD) has emerged as a potential strategy to collect additional evidence on the effectiveness of products after launch. HTA bodies are in the process of adopting real-world evidence (RWE), evidence generated from RWD, into their decision-making processes and setting standards for its use.

As NICE develops best practices on real-world evidence (RWE) use, it is exploring RWD comparative effectiveness studies as a way to enhance its decision-making and address uncertainties in its assessments and guidelines.<sup>1</sup> As part of NICE's five-year strategic vision,<sup>1</sup> NICE is initiating RWE research projects to fill known evidence gaps, to inform when and how RWE can be used in its decision-making, and to determine when after drug launch sample size is sufficient for comparative effectiveness analysis and if the results are consistent over time. NICE is interested in a proof-of-concept RWE comparative effectiveness study to inform these best practices.

# <span id="page-19-1"></span>*5.1. Aetion's Research Collaboration with NICE*

Aetion entered a research collaboration agreement (RCA) with NICE in 2021 to evaluate how RWE studies can be used to fill evidence gaps and reduce uncertainties in NICE assessment and guideline development. Aetion is collaborating with NICE to identify research questions relevant to NICE's RWE standards workstreams. Using the comparative effectiveness study described below, Aetion and NICE are collaborating on identifying fit-for-purpose real-world data, developing the protocol, executing the study using the Aetion Evidence Platform® (2021), software for real-world data analysis (Aetion, Inc. https://www.aetion.com.), and piloting the Structured Template and Reporting Tool for Real World Evidence (STaRT-RWE)<sup>2</sup> for transparent reporting of study implementation.

# <span id="page-19-2"></span>*5.2. Selection of Atrial Fibrillation as Proof of Concept Study*

The safety and efficacy of direct oral anticoagulants (DOACs) compared to vitamin K antagonists (VKAs) in patients with atrial fibrillation (AFib) have been established in randomized clinical trials.<sup>3</sup> In Europe, the following DOACs are approved to prevent venous thromboembolism, deep vein thrombosis, pulmonary embolism, and stroke in adults with AFib: Apixaban (Eliquis®, first authorised in the EU and the UK May 2011); dabigatran (Pradaxa®, March 2008); rivaroxaban (Xarelto®, September 2008); and edoxaban (Lixiana®, June 2015).<sup>4,5</sup>

Between March 2012 and September 2015, NICE separately assessed and recommended four DOACs: dabigatran, rivaroxaban, apixaban, and edoxaban. Each submitted clinical trial compared the investigational DOACs to warfarin; however, there were no head-to-head clinical trials comparing the DOACs to each other at the time of NICE's assessments. Therefore, a network meta-analysis was completed for indirect comparison of individual DOACs agents using warfarin as the common comparator. NICE interpreted the network meta-analysis findings with caution noting the results were not sufficiently robust to differentiate between the products' effectiveness. These uncertainties were reiterated in NICE's 2014 AFib management guidelines in 2014 and 2021 updates. $6,7$ 

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The lack of relevant direct comparisons driving uncertainty is common at the time of NICE assessment, and these uncertainties are not unique to AFib. The uncertainties in comparative effectiveness of AFib treatments impact downstream decisions by healthcare providers and patients as they try to identify the most effective, safe, and cost-effective treatment option for stroke prevention.

NICE has selected AFib for the proof-of-concept comparative effectiveness real-world effectiveness (RWE) study in part because there is a growing body of literature that has successfully validated the use of real-world data (RWD) in  $AFib$ ,<sup>8,9</sup> and the burden of disease and cost to the healthcare system associated with AFib in the UK.<sup>7</sup> In addition, the DOACs for nonvalvular AFib included in this study were approved more than 5 years ago in the UK (edoxaban was approved in June  $2015<sup>10</sup>$ ), which allows us to explore how the value of RWE comparative effectiveness changes with length of follow-up. This can help inform optimal timing for reassessments.

<span id="page-20-0"></span>This study will not only inform when and how NICE could use RWE to address uncertainties in the clinical evidence, but it could also inform NICE's AFib clinical guidelines.

# *5.3. DOACs Therapies for Atrial Fibrillation*

Atrial fibrillation (AFib) is the most common type of cardiac conduction disorder accounting for up to 2.6% of healthcare costs in Europe, and about 9.5% of the European population is estimated to have AFib by 2060.<sup>11</sup> In the UK, AFib affects 1.4 million patients,<sup>12</sup> and between 0.9% to 1.6% of UK's National Health Service (NHS) spendings are attributable to AFib predominately from hospitalizations.<sup>13</sup> The condition is associated with significant complications, including stroke. Nonvalvular AFib increases an individual's risk of stroke by about five times,<sup>14</sup> and between 20% and 30% of stroke cases are attributed to AFib.<sup>11</sup>

Anticoagulants, including VKAs and DOACs, are highly effective in the prevention and treatment of thromboembolic events associated with AFib. Nevertheless, VKAs like warfarin are narrow therapeutic index drugs requiring frequent coagulation monitoring, and have multiple drug and food interactions. Alternatively, DOACs inhibit coagulation via direct and specific binding to active sites of thrombin (e.g., dabigatran) or factor Xa (e.g., apixaban, rivaroxaban, and edoxaban) of the coagulation pathway. Compared to VKAs, DOACs have a wider therapeutic index, which permits use in fixed doses without coagulation monitoring, and relatively limited drug and food interactions. Data from direct comparisons of DOACs (e.g., apixaban vs. dabigatran) are not available, and relative safety and efficacy findings are based on indirect comparisons from network meta-analyses.15,16 Healthcare providers face challenges from uncertainties around DOACs treatment decisions for patients with AFib, particularly around which individual agent has the best benefit-risk profile within the class for mortality and cardiovascular risk reduction.

The current protocol details a study designed to evaluate the effectiveness and safety of the DOACs of interest (apixaban, rivaroxaban, edoxaban, and dabigatran) in direct comparisons, among adults with nonvalvular AFib at risk for stroke in the UK.

# <span id="page-20-1"></span>**6. Research Objectives**

This study evaluates the comparative effectiveness of apixaban versus rivaroxaban (primary objective) for reducing stroke among patients with nonvalvular AFib in the UK.

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<span id="page-21-0"></span>Secondary objectives included comparing the effectiveness of apixaban and rivaroxaban to the other DOACs (edoxaban, and dabigatran).

### $6.1.$ **Hypothetical Target Trial**

When planning and designing RWE studies, it is important to consider emulating a hypothetical target trial.<sup>17</sup> We sought to emulate a four-arm randomized trial that randomizes patients to apixaban, rivaroxaban, edoxaban, or dabigatran. The objective of the hypothetical trial is to compare apixaban to rivaroxaban, as the primary comparison, in the reduction of stroke among patients with nonvalvular AFib in the UK (Figure 6.1). The hypothetical target trial diagram is based on four clinical trials that evaluated the efficacy of respective DOACs in comparison with warfarin in patients with AFib (ARISTOTLE trial studied apixaban;<sup>18</sup> ROCKET AF studied rivaroxaban;<sup>19</sup> ENGAGE AF-TIMI 48 trial studied edoxaban;<sup>20</sup> and RE-LY trial studied dabigatran).<sup>21</sup> In addition, comparisons of individual DOACs to other DOACs as a pharmacological class will be included as secondary comparisons. The hypothetical target trial would be powered for the primary comparison of apixaban versus rivaroxaban, and the study designed to meet diagnostic criteria for the main analysis.

Within the UK, apixaban and rivaroxaban are the most frequently prescribed DOACs (Figure 6.2). Secondary comparisons will include comparing apixaban and rivaroxaban to the two other DOACs (edoxaban and dabigatran) (Table 6.1).



Figure 6.1 **Hypothetical Target Trial Design** 

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### **Figure 6.2 Utilization of DOACs in the UK (April 2016 - January 2021; Source: OpenPrescribing)**

|                    | Rivaroxaban | Edoxaban   | Dabigatran   | <b>DOACs Class</b>                                     |
|--------------------|-------------|--|--|--|
| Apixaban           | Primary     | Secondary  | Secondary  | Secondary<br>(rivaroxaban,<br>edoxaban.<br>dabigatran) |
| Rivaroxaban        |             | Secondary  | Secondary  | Secondary<br>(apixaban,<br>edoxaban.<br>dabigatran)    |
| <b>DOACs Class</b> |             | Secondary<br>(apixaban,<br>rivaroxaban,<br>dabigatran) | Secondary<br>(apixaban,<br>rivaroxaban,<br>edoxaban) |  |

**Table 6.1 Direct Comparisons of Study Exposures of Interest**

# <span id="page-22-0"></span>*6.2. Primary Objective*

The primary objective is to estimate the incidence rates and analogous hazard ratio of stroke (ischemic or hemorrhagic) for patients with nonvalvular AFib who initiated apixaban compared to rivaroxaban (see Table 6.1).

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# <span id="page-23-0"></span>*6.3. Secondary Objectives*

The secondary objectives are to:

- Estimate the incidence rates and analogous hazard ratios of stroke (ischemic or hemorrhagic) for patients with nonvalvular AFib who initiated:
	- apixaban compared to edoxaban, dabigatran, and DOACs class
	- rivaroxaban compared to edoxaban, dabigatran, and DOACs class
	- edoxaban compared to DOACs class
	- dabigatran compared to DOACs class
- Estimate the incidence rates and evaluate the analogous hazard ratio of all-cause mortality, myocardial infarction (MI), transient ischemic attack (TIA), major bleeding events, and major adverse cardiovascular events (MACE) for patients with nonvalvular AFib who initiated individual DOACs compared to those who initiated other DOACs (see Table 6.1).

# <span id="page-23-1"></span>**7. Data Source**

This study uses anonymized patient RWD from The Health Improvement Network (THIN®) Database (A Cegedim Proprietary Database). THIN® is an unobtrusive medical data collection scheme that collects anonymised patient data from its members. THIN® data is a primary care research database containing anonymised electronic health record data from around 850 UK general practices (GPs) using the VISION clinical system (since 1994), containing records around 20 million patients, of which 2.6 million are currently active. THIN® collates the computerized medical records of general practitioners (GPs) in the UK who act as the gatekeepers of healthcare and maintain patients' life-long electronic health records. As such they are responsible for primary healthcare and specialist referrals, and they also store information stemming from specialist referrals, and hospitalizations. GPs act as the first point of contact for any non-emergency health-related issues, which may then be managed within primary care and/or referred to secondary care as necessary. The THIN® database includes demographic, medical diagnosis, patient medical history, lifestyle factors, and written prescription information on individual patients. The Quality and Outcome Framework (QOF) rewards practices for the quality care they provide their patients, based on a number of indicators, including medical record data entry.

The THIN® database has been used for scientific publications in AF and stroke research<sup>22-25</sup> and has been validated and widely used in pharmacoepidemiologic studies.<sup>22</sup>

# <span id="page-23-2"></span>**8. Research Methods**

# <span id="page-23-3"></span>*8.1. Study Design*

This study is a non-interventional, retrospective cohort study utilising RWD from the UK to address the study objectives. Patients with nonvalvular AFib who initiated (i.e., new users of) DOACs during the observation period will comprise the study cohort. The exposure groups of interest will be patients initiating individual DOACs (apixaban, rivaroxaban, edoxaban, and dabigatran). All patients will be followed for the first occurrence of each of the study outcomes of interest.

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# <span id="page-24-0"></span>*8.2. Setting*

<span id="page-24-1"></span>This study will be conducted using an electronic anonymized patient RWD from The Health Improvement Network (THIN®) Database (A Cegedim Proprietary Database) that includes data collected in primary care settings in the UK.

# **8.2.1. Study Population**

This study will be conducted using patient clinical data extracted from the THIN® database in the UK for adults with nonvalvular AFib at risk for stroke who initiated DOACs between July 1, 2014 and December 31, 2020. This period was selected to minimize potential effects on recommended treatment options due to changes in AFib management guidelines. The NICE AFib management guidelines, which recommend prescribing a DOACs (apixaban, rivaroxaban, edoxaban, or dabigatran) for stroke prevention were published in June  $2014^6$  and were not updated until April 2021<sup>7</sup>, thus we believe treatment options to be stable during this period. Patients with AFib are identified by Read Medical Codes and ICD-10 diagnosis codes.

<span id="page-24-2"></span>Patient selection criteria are described in the following sections and Figure 8.1.

# **8.2.1.1. Inclusion Criteria**

The following are the inclusion criteria for the study cohort:

<span id="page-24-3"></span>● ≥1 prescription for DOACs of interest (index date)

# **8.2.1.2. Exclusion Criteria**

The following are the exclusion criteria for the study cohort:

- Exclusion criteria applied during selection of the data cut:
	- <1 diagnosis code for AFib on index date or in 12 months prior (applied during selection of the data cut)
	- $\circ$  <1 risk factor for stroke other than Afib on index date or in 12 months prior defined by  $CHA<sub>2</sub>DS<sub>2</sub> VASC > 1$  for men and  $>2$  for women<sup>26</sup>
	- <1 medical encounter in the 180 days prior to index date
	- $\circ$   $\geq$  1 diagnosis code for the following conditions on index date or any time prior:
		- Cardiac valve disease (mitral stenosis, valvular replacement)
			- Deep vein thrombosis
			- Pulmonary embolism
	- Exclusion criteria applied after selection of the data cut:
		- <18 years of age on index date
			- <12 months' registration with a GP prior to the index date
			- $\circ$   $\geq$  1 prescription record for DOACs in the 12 months prior to the index date (i.e., prevalent users)
			- ≥1 diagnosis code for the study outcomes of interest (stroke, MI, TIA, major bleeding events) on index date or any time prior
			- ≥1 diagnosis code for the following conditions on index date or any time prior: **Angina** 
				-
				- Congenital heart disease
			- Missing age
			- Missing gender

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# <span id="page-25-0"></span>**8.2.2. Study Periods**

The study observation period spans July 1, 2013 through December 31, 2020, consisting of a drug indexing, baseline and follow-up periods (Figure 8.1). The drug indexing period will be used to ascertain DOACs initiation (index date). Baseline period is defined as 12 months prior to and including the index date. The follow-up period will include all available time after study index date. Patients will be followed in an "intention to treat" (ITT) approach from their index date until the end of follow-up period or the occurrence of the following events, whichever occurs first:

- First occurrence of the study outcomes of interest
- Death
- End of patient data
- End of patient registration with GP.

In addition, patients meeting the selection criteria will be followed in an "as-treated" approach from their index date until the end of the study follow-up period or the occurrence of the following events, whichever occurs first:

- First occurrence of the study outcomes of interest.
- Death
- End of patient data.
- End of patient registration with GP.
- Termination of exposure, crossover of exposure group, or addition of another DOAC.

Termination of exposure in the as-treated analysis is defined as having a gap exceeding 30 days (grace period) between the end of a prescription, based on its start date and duration, and the start of the next prescription, or if no additional prescription occurs. The termination date is defined as the prescription end date plus a 30-day risk window.

![](_page_26_Figure_4.jpeg)

Figure 8.1 **Study Design Diagram** 

### <span id="page-26-0"></span> $8.3.$ **Variables**

The following sections describe the study exposures, outcomes, and covariates (including patient demographics, clinical characteristics, comorbidities, and concomitant medications).

# <span id="page-26-1"></span>8.3.1. Exposures

DOACs of interest will include apixaban, rivaroxaban, edoxaban, and dabigatran. Exposures will be defined by corresponding ATC codes: B01AF02; B01AF01; B01AF03; and B01AE07, respectively. Initiators of DOACs are defined as new users of DOACs with no record of prescription for any DOAC during 12 months before index date.

Only a single index date is allowed per patient; if a patient qualifies as a new user of an exposure group multiple times, the patient will enter the cohort on the first qualifying exposure date. Patients with exposure to more than one DOAC on the index date will be excluded. Duration of treatment is calculated from the provided days supplied/duration data field. Fields with 0 are assumed to be 1 day. A grace period of 30 days between refills and risk window of 30 days after the last refill is used to define persistence. The operationalized definition of exposure will be recorded in the STaRT RWE template and is available on request.

# <span id="page-26-2"></span>8.3.1.1. Primary Exposure Groups

- *Exposure Group 1:* Initiators of apixaban
- *Exposure Group 2:* Initiators of rivaroxaban  $\bullet$

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# <span id="page-27-0"></span>**8.3.1.2. Secondary Exposure Groups**

The following independent Exposure Groups will be created:

- *Exposure Group 3:* Initiators of edoxaban
- *Exposure Group 4:* Initiators of dabigatran
- *Exposure Group 5:* Initiators of rivaroxaban, edoxaban, and dabigatran
- *Exposure Group 6:* Initiators of apixaban, edoxaban, and dabigatran
- <span id="page-27-1"></span>*● Exposure Group 7:* Initiators of apixaban, rivaroxaban, and dabigatran
- *Exposure Group 8:* Initiators of apixaban, rivaroxaban, and edoxaban

# <span id="page-27-2"></span>**8.3.1.3. Primary Comparison**

● *Primary Comparison:* initiators of apixaban vs initiators of rivaroxaban (i.e., *Exposure Group 1* vs *Exposure Group 2*)

# **8.3.1.4. Secondary Comparisons**

- *Secondary Comparison 1:* initiators of apixaban vs initiators of edoxaban (i.e., *Exposure Group 1* vs *Exposure Group 3*)
- *Secondary Comparison 2:* initiators of apixaban vs initiators of dabigatran (i.e., *Exposure Group 1* vs *Exposure Group 4*)
- *Secondary Comparison 3:* initiators of apixaban vs initiators of other DOACs class (i.e., *Exposure Group 1* vs *Exposure Group 5*)
- *Secondary Comparison 4:* initiators of rivaroxaban vs initiators of edoxaban (i.e., *Exposure Group 2* vs *Exposure Group 3*)
- *Secondary Comparison 5:* initiators of rivaroxaban vs initiators of dabigatran (i.e., *Exposure Group 2* vs *Exposure Group 4*)
- *Secondary Comparison 6:* initiators of rivaroxaban vs initiators of other DOACs class (i.e., *Exposure Group 2* vs *Exposure Group 6*)
- *●*
- *Secondary Comparison 7:* Initiators of edoxaban vs initiators of other DOACs class (i.e., *Exposure Group 3* vs *Exposure Group 7*)
- *Secondary Comparison 8:* initiators of dabigatran vs initiators of other DOACs class (i.e., *Exposure Group 4* vs *Exposure Group 8*)

# <span id="page-27-3"></span>**8.3.2. Outcomes**

The primary outcome is ischemic or hemorrhagic stroke. Secondary outcomes include the occurrence of all-cause mortality, myocardial infarction, transient ischemic attacks, major bleeding events, and major adverse cardiovascular events (MACE). Major bleeding is defined as a composite outcome of major intracranial (including hemorrhagic stroke), gastrointestinal, and urogenital bleeds. MACE outcomes will be a composite endpoint of angina, myocardial infarction, and stroke.

Outcomes of interest will be measured during the study follow-up period, and defined by corresponding Read Medical Codes and ICD-10 diagnosis codes at the primary care setting. The operational definitions of each outcome will be recorded in the STaRT-RWE template and are available on request.

# **8.3.3. Covariates**

<span id="page-27-4"></span>Patient demographics, clinical characteristics, comorbidities, and concomitant medications will be measured during baseline period and index date, and identified by Read Medical Codes,

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ICD-10 diagnosis codes, and medication ATC codes for corresponding covariates. These covariates include cohort entry year, age, sex, socioeconomic indicators (e.g., marital status), and selected components of  $CHA<sub>2</sub>DS<sub>2</sub> VASC$  stroke risk score<sup>27,28</sup> (heart failure, hypertension, and diabetes), and ORBIT-AF bleeding risk score in AFib patients<sup>29,30</sup> (heart failure, renal impairment, hepatic impairment; osteoporosis/hip fractures, anemia, antiplatelet therapy, aspirin therapy, NSAIDs therapy, and smoking and alcohol drinking history [if available]).

<span id="page-28-0"></span>Additional comorbidities and concomitant medications deemed to be associated with the primary and secondary outcomes of interest will be included in the analyses. All covariates will be determined based on the literature and clinical knowledge. The operational definitions of each covariate will be recorded in the STaRT-RWE template and are available on request.

# 9. Data Analysis

Data analyses will be performed in a stepwise approach involving three distinct sequential phases (Figure 9.1): *Exploratory Phase* (data explorations will be done in advance to inform key design decisions), Diagnostic Phase (requirements that must be met prior to viewing study outcomes, e.g. covariate balance), and Inferential Phase (comparative analyses).

![](_page_28_Figure_8.jpeg)

**Figure 9.1 Data Analysis Phases** 

In the Exploratory Phase, the study implementation team completed explorations related to the patient selection criteria, exposure groups, and outcome definitions that informed the development of the study protocol.

For the Diagnostic Phase, the checklist in Table 9.1 must be satisfactorily completed prior to beginning the implementation of the Inferential Phase. As such, the relationship between the exposure and outcome of interest will not be described or evaluated in the analytic dataset until the study implementation team reaches consensus that the diagnostic criteria are satisfied.

![](_page_28_Picture_99.jpeg)

![](_page_28_Picture_100.jpeg)

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![](_page_29_Picture_187.jpeg)

Unless otherwise specified, results will be provided as descriptive statistics with categorical variables reported using frequency distributions, and continuous variables reported using means, standard deviations, medians, minimums, maximums, 25th percentiles and 75th

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percentiles. For inferential statistics, both crude and covariate adjusted analyses will be reported for study outcomes.

Among patients in each of the exposure groups, the incidence rates and 95% confidence intervals (CI) of stroke and secondary outcomes will be calculated for each of the mutually exclusive comparison groups, along with the number of events, total number of individuals, and accrued person-time. Incidence rates will be calculated as the number of incident outcomes of interest within the follow-up period divided by the total person-time at-risk, and reported as rate per 1,000 person-years.

In each comparison cohort, multivariable logistic regression will be used to estimate the probability of a patient's initiation of the exposure given baseline covariates and cohort entry year (i.e., the propensity score). Propensity score matching between exposure groups will be performed using 1:1 nearest neighbor matching without replacement with a maximum matching caliper of 0.01. In addition to graphical depictions of propensity score distributions, the absolute standardized differences (ASD) in proportions and means of baseline characteristics will be estimated to examine comparability of exposure groups. Covariates with ASD ≤0.10 are considered balanced between comparator groups.<sup>33</sup> Covariates with small imbalances  $0.10 \le$  $ASD \leq 0.15$ ) may be deemed balanced if the covariate does not predict the outcome among the referent group.

Cox proportional hazards regression (outcomes model) will be used to estimate hazard ratios (HR) and 95% CI for each outcome after propensity score matching. The incidence of stroke and secondary outcomes will be compared between individual DOACs in primary and secondary comparisons as mutually exclusive cohorts.

Subgroup analyses that will be considered include DOACs initiators with prior warfarin use; age  $($  <75 and >75 years); concomitant aspirin use;  $CHA<sub>z</sub>DS<sub>z</sub>$  VASc score (0-1, 2-3, and >4), and selected comorbid conditions, e.g. diabetes, heart failure, and BMI status (<30 and  $\geq$ 30 kg/m<sup>2</sup>).

Sensitivity analyses will be performed to assess the impact of varying the study period on study findings by limiting analyses on data up to February 28, 2020 to account for potential impact of COVID-19 pandemic on healthcare utilization.

Additionally, high-dimensional propensity score (HdPS) analysis will be used as a sensitivity analysis to estimate the association between treatment with DOACs and the primary outcome of stroke. The HdPS approach is a 7-step algorithm that empirically identifies a pool of covariates from different data dimensions based on their prevalence, and then selects a subset of the covariates for inclusion in a propensity score model based on their potential to bias the exposure-outcome association.<sup>30</sup>

Missing values will be reported as missing, and no imputation will be undertaken. All data analysis will be conducted using Aetion Evidence Platform® (2021), software for real-world data analysis. Aetion, Inc. https://www.aetion.com.

# <span id="page-30-0"></span>**10. Limitations of the Research Methods**

This study will be based on secondary use data from EHR with inherent limitations, including misclassification. For example, a patient may not necessarily use a medication that they have received a prescription for; however, this is less likely in the case of medications for chronic conditions like cardiovascular diseases. Additionally, while it is possible to identify study

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outcomes of interest in the THIN® database, misclassification of the outcome is possible, as the presence of a diagnosis in medical records does not necessarily indicate a true presence of disease; however, the outcomes of interest are included in the Quality and Outcome Framework  $(QOF)^{34}$  which rewards and incentivizes GPs in England for quality care and coding of diagnoses. In addition, validated algorithms from previously published real-world studies will be used to define study variables of interest.

An incident user design will be applied to reduce biases that can impact non-randomized studies, especially when using EHR.<sup>35</sup> An ITT analysis will be used because it has the advantage of eliminating certain types of biases by preserving the prognostic balance obtained through the propensity score matching and maintaining sample size;<sup>36,37</sup> however, ITT analysis is sometimes considered to be "too conservative" and has the possibility of exposure misclassification.35,36 Additionally, "as-treated" approach will be used that will censor individuals upon discontinuation of index treatment or treatment crossover. However, limitations of an "as-treated" analysis include the possibility of introducing certain biases including differential and informative censoring, or time-dependent confounding biases. $36,37$ 

Finally, while the THIN® database contains substantial information for inclusion/exclusion criteria and confounder control, medical conditions or a family history of medical conditions are only ascertainable where established diagnoses and procedures for those conditions exist. Additionally, some key covariates may not be available (unmeasured or imperfectly measured confounders), e.g. alcohol and smoking histories are not routinely recorded. Residual confounding by unmeasured variables will likely be present; however, efforts will be made to incorporate all potential confounding variables into the propensity score analysis to minimize impacts from unmeasured confounding, including the application of HdPS sensitivity analysis. Compared to conventional confounding adjustment methods, HdPS algorithm improves confounding control in situations when the variables are weak confounders that are weakly associated with exposure and themselves are associated with unmeasured confounders.<sup>38</sup>

# <span id="page-31-0"></span>**11. Quality Control**

The study data to be used for the proposed analysis will be connected to the Aetion Evidence Platform (AEP), which will be used to perform all the analyses. The AEP is a data-handling technology, which allows for the analysis of large patient claims, EHR, and other transactional datasets by indexing patient data into a form that can be queried by an internal patient variable language. Data is minimally transformed at the point of connection to the AEP, thus the original format of the THIN® data is preserved. At the point of data connection to the platform some discard rules are applied. Patient events are excluded if there are no dates associated with them, or if the start date of the event is preceded by the end date of the event (e.g. discharge date precedes admission date for an inpatient event). Aetion IDs are assigned to THIN® patient IDs and a crosswalk file is kept as a protected file available upon request to authorized parties. The patient data is individual level patient data and will be analyzed within the AEP. Aggregated results will be exported from the platform in the form of tables.

THIN® data are loaded into the AEP after minimal processing into patient longitudinal timelines to enable representation of the original data and without any data loss. The following data checks are performed during the data connection process:

- Events are required to have a valid start date.
- Record counts are cross-checked for validation and compared to the original data counts.

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● Enrollment information is processed and checked, and all fields and sample records are checked to ensure validity.

The process is implemented via a double-programming, rule-based approach that is flexible, automated, and scalable so that it can be reproduced when data is updated. Data processed through the data connector (coded by the Aetion Data Engineering team) are compared against data transformed independently to reproduce the data connector logic (coded by the Aetion Data Science team). Checks are performed first on a subset of data, then on the full data. This process ensures that customers are working with a scientifically valid data product when they perform analyses on the data using the AEP. Statistical analyses will be conducted using the validated AEP.

The Aetion Science Team will build measures for cohort inclusion/exclusion criteria, outcomes and covariates. All measures created, cohorts developed, statistical analyses implemented, and tables populated will undergo quality control review by at least one additional analyst or scientist under the supervision of the Senior Scientist. Quality control methods include checks for the validity and logical content of codes and checks for missing values and variables. In order to control for potential inconsistencies and errors, all variables will be tabulated. In addition, the distribution of values for each variable, including potential outliers, will be examined. This protocol will be strictly followed in the study implementation. However, variable definitions may undergo modification if determined to be scientifically sensible. All changes to this protocol will be documented in protocol amendments.

# <span id="page-32-0"></span>**12. Protection of Human Subjects**

This study will be conducted in accordance with applicable laws and regulations of the country where the study is being conducted, as appropriate. This observational, non-interventional study does not affect the treatment of the patients. The study is conducted in accordance with ethical principles that are consistent with the Declaration of Helsinki, International Conference on Harmonisation (ICH) Good Clinical Practice (GCP), Good Pharmacoepidemiology Practice (GPP), and the applicable legislation on non-interventional studies and/or observational studies.

The study concept was approved by the THIN® Scientific Review Committee on July 6, 2021.

# <span id="page-32-1"></span>**13. Management and Reporting of Adverse Events/Adverse Reactions**

This is a non-interventional study using only structured secondary data, and attribution of adverse reactions/adverse events to specific exposure is not possible, and therefore no individual case safety report (ICSR) reporting is required. During the course of observational research using existing secondary databases, the proposed study will use structured data fields only and will not involve chart review or validation to obtain additional information on the adverse events other than the study outcomes of interest.

# <span id="page-32-2"></span>**14. Plans for Disseminating and Communicating Study Results**

This study protocol will be registered in the ENCePP EU PAS Register<sup>39</sup> prior to the implementation of the diagnostic phase. In addition, study findings will be disseminated as manuscript(s) in peer-reviewed journals and/or as conference abstract presentations at international professional conferences.

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# **Aetion Measure Appendix**

![](_page_36_Picture_203.jpeg)

![](_page_36_Picture_204.jpeg)

![](_page_36_Picture_205.jpeg)

![](_page_36_Picture_206.jpeg)

![](_page_36_Picture_207.jpeg)

![](_page_36_Picture_208.jpeg)

**Human Subjects/Ethics Approval**

This study will be conducted in accordance with applicable laws and regulations of the country where the study is being conducted, as appropriate. This observational, non-interventional study does not affect the treatment of the patients. The study is conducted in accordance with ethical principles that are  $\epsilon$ 

*This Measure Appendix is based largely on the STaRT-RWE templates proposed by Wang et. al., 2020.* 

### **Research Question Overview**

![](_page_38_Figure_4.jpeg)

![](_page_38_Figure_5.jpeg)

![](_page_39_Picture_40.jpeg)

![](_page_40_Picture_82.jpeg)

**Index** 

![](_page_41_Picture_465.jpeg)

The criterion that define the date of entry to the cohon(s) is specified in this section. There should be one row for each unique definition of a study population entry. If the study is elecriptive, there may only be one m

Check the pre-specified box if the exclusion criterion was specified before beginning data analyses, check the varied for sensitivity box if it was modified as part of sensitivity analyses. Specify the source of algorithms

![](_page_42_Picture_511.jpeg)

### **Covariate Drugs**

![](_page_43_Picture_365.jpeg)

![](_page_44_Picture_343.jpeg)

![](_page_44_Picture_344.jpeg)

**STaRT-RWE Instructions:** 

Define the covariate conceptually, with accompanying details as necessary. Specify which planned analyses adjust for the covariate, and how it is specified in the analysis (e.g., continuous, categorical, binary). D assessment window relative to the index date (day 0), whether there are restrictions on care setting or diagnosis position in the algorithm, and which study populations defined in Table 3B the covariate is measure Specify the source of algorithms to define covariates.

Check the pre-specified box if the covariate was specified before beginning data analyses, check the varied for sensitivity box if it was modified as part of sensitivity analyses. Specify the source of algorithms to d covariates.

### **Covariate Conditions**

![](_page_45_Picture_471.jpeg)

![](_page_46_Picture_138.jpeg)

### **STaRT-RWE Instructions:**

Specify when follow up begins relative to the index date (day 0) and select each criterion that is used to end follow up.

Check the pre-specified box if the outcome parameters were specified before beginning data analyses, check the varied for sensitivity box if the parameters were modified as part of sensitivity analyses.

### **Outcomes**

![](_page_47_Picture_372.jpeg)

![](_page_48_Picture_229.jpeg)

![](_page_49_Picture_146.jpeg)

*This Measure Appendix is based largely on the START-RWE templates proposed by Wang et. al., 2020.* 

![](_page_50_Picture_317.jpeg)

![](_page_51_Picture_182.jpeg)

![](_page_52_Picture_144.jpeg)

![](_page_52_Picture_145.jpeg)

![](_page_52_Picture_146.jpeg)

![](_page_52_Picture_147.jpeg)

![](_page_52_Picture_148.jpeg)

![](_page_53_Picture_303.jpeg)

![](_page_54_Picture_252.jpeg)

![](_page_55_Picture_407.jpeg)

![](_page_56_Picture_351.jpeg)

![](_page_57_Picture_247.jpeg)

![](_page_58_Picture_340.jpeg)

![](_page_59_Picture_185.jpeg)

![](_page_60_Picture_329.jpeg)

![](_page_61_Picture_283.jpeg)

![](_page_62_Picture_228.jpeg)

Supplemental Table 1. Baseline characteristics of patients with nonvalvular atrial fibrillation newly initiating a DOAC, before and after propensity score matching: characteristics not included in propensity score models d

![](_page_63_Picture_997.jpeg)

Characteristics assessed in the year prior to and including index date.<br>SD = standard deviation; IQR = interquartile range; DOAC = direct oral anticoagulants; ASD = absolute standardized difference.

\* Values reported reflect all available data without modification or cleaning.

\*\* Patients with prescriptions for multiple dosages are reported as having all dosages observed

ouppiemental rabie z. Per-protocol analysis using a obou baseline period for an characteristics: baseline characteristics of patents with nonvarythar atrial infinition newly infidating a DOAG, before and after propensity s **matching**

![](_page_64_Picture_1331.jpeg)

![](_page_65_Picture_443.jpeg)

Characteristics reported as N (%) unless otherwise specified.

SD = standard deviation; IQR = interquartile range; DOAC = direct oral anticoagulants; ASD = absolute standardized difference.

**Supplemental Table 3. Per-protocol analysis using a 365d baseline period for all characteristics: Hazard ratio of stroke and secondary outcomes among patients with nonvalvular atrial fibrillation newly initiating DOACs after propensity score matching**

![](_page_66_Picture_530.jpeg)

### **Apixaban vs rivaroxaban: primary outcome (stroke) among subgroups, ITT**

![](_page_66_Picture_531.jpeg)

![](_page_66_Picture_532.jpeg)

PY = person-years; aHR = adjusted hazard ratio; CI = confidence interval; PS = propensity score; ITT = intent-to-treat; AT = as-treatec

PS model accounts for age, gender, CHA2DS2 VASc score, year of treatment initiation, and the following diagnoses and treatments in baseline: non-major bleeding events, anemia, diabetes, hypertension, heart failure, osteoporosis/hip fracture, malignant neoplasm, acute kidney injury, chronic kidney disease, asthma/copd, dementia, aspirin, antiplatelets other than aspirin, warfarin, antiameric preparations, NSAIDs, opioids, SSRIs, antidepressants other than SSRIs, antiepileptics, antipsychotics, benzodiazepines, lipid

In analysis of stroke, MI, TIA, major bleeding events, and AMS, patients were followed until occurrence of outcome, death, end of patient registration, or end of study period (12/2020).

\* In analysis of all cause mortality, patients were followed until occurrence of outcome (death), end of study period (12/2020), or the later date of end of patient registration and any recorded death within 90 days of end of patient registration. Propensity score matched sample size for analysis of all cause mortality differs from sample size in analysis of other outcomes because of differences in censoring criteria which impact a small number of patients' eligibility for inclusion in analysis at the start of follow-up. DOACs comprised apixaban, rivaroxaban, edoxaban, dabigatran

![](_page_67_Picture_257.jpeg)

### **Supplemental Table 4: Length of follow-up and censoring reasons for the primary comparison (stroke, ITT and AT analyses) and secondary comparisons (stroke, ITT analyses) after PS-matching**

DOAC = direct oral anticoagulant; ITT = intent-to-treat; AT = as-treated; PS = propensity score.

Apixaban was compared with other DOACs which included rivaroxaban, edoxaban, and dabigatran; rivaroxaban was compared with other DOACs which included apixaban, edoxaban, and dabigatran.